# Consistent and computationally efficient estimation for stochastic LPV state-space models: realization based approach.

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#### Outline

#### Contributions

- Realization and identification problem for stochastic LPV-SSA representations.
- stochastic covariance realization algorithm.
- Identification algorithm
- Simulation case study

LPV State-Space Affine (LPV-SSA) representations:

$$\mathbf{x}(t+1) = A(\boldsymbol{\mu}(t))\mathbf{x}(t) + \mathcal{K}(\boldsymbol{\mu}(t))\mathbf{v}(t),$$
  
 $\bar{\mathbf{y}}(t) = C\mathbf{x}(t) + D\mathbf{v}(t)$ 

- $(x, \bar{y}, v)$  denote state, output, noise *processes*,
- $ightharpoonup \mu$  scheduling signal process taking values in  $\mathbb{R}^d$ .

$$A(\mu(t)) = A_1 + \sum_{i=2}^d A_i \mu_i(t), \quad K(\mu(t)) = K_1 + \sum_{i=2}^d K_i \mu_i(t).$$

Noise covariance  $Q_i = \mathbb{E}\left[\mathbf{v}(t)\mathbf{v}^{\top}(t)\mu_i(t)^2\right] \quad \forall i = 1, \cdots, d$ 

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## Problem (Realization problem)

For observed process  $(\mathbf{y}, \boldsymbol{\mu})$ , find matrices  $(\{A_i, K_i\}_{i=1}^d, C, D)$  and processes  $\mathbf{x}, \mathbf{v}$  such that LPV-SSA is a realization of  $\mathbf{y}$ , i.e.,  $\mathbf{y}(t) = \bar{\mathbf{y}}(t), \ t \in \mathbb{Z}$ .

#### Problem (Identification problem)

Given sample paths  $\{y(t), \mu(t)\}_{t=1}^N$ , compute the estimates  $\{\{\hat{A}_i^N, \hat{K}_i^N, \hat{Q}_i^N\}_{i=1}^d, \hat{C}^N, \hat{D}^N\}$ , such that as  $N \to \infty$ , the estimated matrices converge to matrices  $\{\{A_i, K_i, Q_i\}_{i=1}^d, C, D\}$  of true system, which is a representation of  $(\mathbf{y}, \boldsymbol{\mu})$ .

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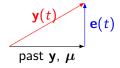
## Assumptions

- The scheduling process  $\boldsymbol{\mu} = [1, \mu_2, \dots, \mu_d]^{\top}$  is i.i.d.  $\mathbb{E}\left[\boldsymbol{\mu}_i^2\right] = p_i$
- ► [x<sup>T</sup> v<sup>T</sup>] satisfies technical conditions: is a Zero Mean Wide Sense Stationary w.r.t. scheduling Inputs
- ▶ The LPV-SSA representation  $(\{A_i, K_i\}_{i=1}^d, C, D, \mathbf{v})$  satisfies technical condition: it is stationary for  $\boldsymbol{\mu}$

Stationary LPV-SSA: stable, additive white noise, statistical independence conditions for the noise, input and state.

#### Forward innovation form

▶ the forward innovation process  $\mathbf{e}$   $\mathbf{e}(t) = \mathbf{y}(t) - E_l[\mathbf{y}(t) \mid \text{past } \mathbf{y}, \boldsymbol{\mu}], E_l \text{ denotes projection.}$ 



► The stationary LPV-SSA  $(\{A_i, K_i\}_{i=1}^d, C, D, \mathbf{v})$  in forward innovation form if  $\mathbf{v} = \mathbf{e}$  and  $D = I_p$ , *i.e.*,

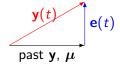
$$\mathbf{x}(t+1) = \sum_{i=1}^{a} \left( A_i \mathbf{x}(t) + K_i \mathbf{e}(t) \right) \boldsymbol{\mu}_i(t), \mathbf{y}(t) = C \mathbf{x}(t) + \mathbf{e}(t).$$

#### Theorem

Under suitable assumptions, every LPV-SSA representation can be transformed to minimal forward innovation form.

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#### **Theorem**

Under suitable assumptions, every LPV-SSA representation can be transformed to minimal forward innovation form.

► The covariance sequence  $\Psi_{\mathbf{y}}(i_1 \cdots i_k)$  is defined as

$$\Psi_{\mathbf{y}}(i_1\cdots i_k) = \mathbb{E}\left[\mathbf{y}(t)\frac{\mu_{i_k}(t-1)}{\sqrt{p_{i_k}}}\cdots \frac{\mu_{i_1}(t-k)}{\sqrt{p_{i_1}}}(\mathbf{y}(t-k))^\top\right] \quad k > 0$$

constructed from past outputs and scheduling.

▶  $\Psi_{\mathbf{y}}(i_1 \cdots i_k)$  are a sub-Markov parameters of deterministic LPV-SSA:  $\Psi_{\mathbf{y}}(i_1 \cdots i_k) = CA_{i_k} \cdots A_{i_1}B_{i_1}$   $z(t+1) = \sum_{d} (A_i z(t) + B_i w(t))\mu_i(t), \ \mathbf{y}(t) = Cz(t)$ 

$$Z(t+1) = \sum_{i=1}^{n} (A_i Z(t) + B_i W(t)) \mu_i(t), \ \mathbf{y}(t) = CZ(t)$$

Let state and noise covariance be  $P_i = \mathbb{E} \left[ \mathbf{x}(t) \mathbf{x}^{\top}(t) \mu_i^2(t) \right], \ Q_i = \mathbb{E} \left[ \mathbf{e}(t) \mathbf{e}^{\top}(t) \mu_i^2(t) \right]$ 

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- $\blacktriangleright B_i = \frac{1}{\sqrt{p_i}} (A_{\sigma} P_i C^{\top} + K_i Q_i) \quad \forall i = 1, \cdots, d$

## Stochastic realization algorithm: Riccati equations

▶ The state and noise covariance satisfy Riccati-like equations

$$P_{i} = \sum_{j} p_{i} \left( A_{j} P_{j} A_{j}^{\top} + K_{j} Q_{j} K_{j}^{\top} \right)$$

$$Q_{i} = p_{i} \mathbb{E} \left[ \mathbf{y}(t) \mathbf{y}^{\top}(t) \boldsymbol{\mu}_{i}^{2}(t) \right] - C P_{i} C^{\top}$$

$$K_{i} = \left( B_{i} \sqrt{p_{i}} - A_{i} P_{i} C^{\top} \right) (Q_{i})^{-1}$$

The equations can be solved iteratively  $k = 1, 2, \dots, \mathcal{I}$ 

$$\begin{split} \hat{P}_i^{k+1} &= \sum_j p_i \left( \hat{A}_j \hat{P}_j^k \hat{A}_j^\top + \hat{K}_j \hat{Q}_j^i \hat{K}_j^\top \right) \\ \hat{Q}_i^k &= p_i \mathbb{E} \left[ \mathbf{y}(t) \mathbf{y}^\top (t) \boldsymbol{\mu}_i^2 (t) \right] - \hat{C} \hat{P}_i^k \hat{C}^\top \\ \hat{K}_i^k &= \left( \hat{B}_i \sqrt{p_i} - \hat{A}_i \hat{P}_i^k (\hat{C})^\top \right) \left( \hat{Q}_i^k \right)^{-1} \end{split}$$

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- 1. Construct covariance sequence  $\Psi_{\mathbf{y}}(i_1 \cdots i_k) = CA_{i_1}A_{i_2} \cdots A_{i_k}B$  from the past  $\mathbf{y}$  and  $\boldsymbol{\mu}$ .
- 2. Construct Hankel matrices  $H_{\Psi_y}$  from the covariance sequence  $\Psi_y(i_1\cdots i_k)$ .
- 3. Estimate  $(\{A_i, B_i\}_{i=1}^d, C)$  from the  $H_{\Psi_y}$ , using Ho-Kalman like realization algorithm.
- 4. Estimate the  $\{K_i\}_{i=1}^d$  and noise-covariances  $\{P_i, Q_i\}_{i=1}^d$  solving Riccatti-like equations.

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## Stochastic realization algorithm: selections

- ▶ The construction of full Hankel matrix  $H_{\Psi_{V}}$  is avoided.
- We use the idea of selections [Cox et. al. '18, Sontag et.al '70] to select specific entries of the Hankel matrix  $H_{\Psi_{\nu}}$
- The selection pair  $(\alpha, \beta)$  is chosen to select the entries of the Hankel matrix  $H_{\Psi_{\nu}}$  to form a sub-Hankel matrix  $H_{\alpha,\beta}$ .
- ▶ The sub-Hankel matrix  $H_{\alpha,\beta}$  is such that,

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## Identification algorithm

Main idea: estimate the covariance sequences  $\Psi_{\mathbf{y}}$  from the observed sample paths  $\{y(t), \mu(t)\}_{t=1}^{N}$ .

- 1. Compute empirical covariances from data, approximating expectations with sample mean
- 2. Run stochastic realization algorithm using empirical covariances for  ${\mathcal I}$  iterations to get
  - ► Model matrices  $(\{\tilde{A}_i, \tilde{K}_i^{\mathcal{I}}\}_{i=1}^d, \tilde{C})$
  - Noise and state covariances  $(\{\tilde{Q}_i^{\mathcal{I}}, \tilde{P}_i^{\mathcal{I}}\}_{i=1}^d)$

## Identification algorithm: Consistency

## Theorem (Consistency)

The estimates satisfy

$$\begin{split} \tilde{K}_i &= \lim_{\mathcal{I} \to \infty} \lim_{N \to \infty} \tilde{K}_i^{N,\mathcal{I}}, \\ \tilde{A}_i &= \lim_{N \to \infty} \tilde{A}_i^N, \\ \tilde{C} &= \lim_{N \to \infty} \tilde{C}^N \end{split}$$

and  $(\{\tilde{A}_i, \tilde{K}_i, \}_{i=1}^d, \tilde{C}, \hat{\mathbf{x}}, \mathbf{e})$  is a minimal stationary LPV-SSA representation of  $(\mathbf{y}, \boldsymbol{\mu})$  and  $\mathbb{E}\left[\mathbf{e}(t)(\mathbf{e}(t))^{\top}\boldsymbol{\mu}_i^2(t)\right] = \lim_{\mathcal{I} \to \infty} \lim_{N \to \infty} \tilde{Q}_i^{N,\mathcal{I}}, i = 1, \cdots d$ 

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## Numerical Example

The LPV-SSA representation in forward innovation form

$$A_1 = \begin{bmatrix} 0.4 & 0.4 & 0 \\ 0 & 0 & 0 \\ 0 & 0 & 0 \end{bmatrix}, A_2 = \begin{bmatrix} 0 & 0 & 0 \\ 0 & 0.4 & 0.4 \\ 0 & 0.4 & 0.4 \end{bmatrix},$$

$$\mathcal{K}_1 = \begin{bmatrix} -0.036 \\ 0 \\ 1 \end{bmatrix}, \mathcal{K}_2 = \begin{bmatrix} 0 \\ 0.015 \\ 1.17 \end{bmatrix}, \mathcal{C} = \begin{bmatrix} 1 & 0 & 0 \end{bmatrix},$$

- ▶ Training data N = 100,000, validation  $N_{\text{val}} = 100000$ .
- Scheduling signal process  $\mu = [\mu_1 \ \mu_2]$  such that  $\mu_1(t) = 1$  and  $\mu_2(t)$  is a white-noise process with  $\mathcal{U}(-1.5, 1.5)$
- ▶ innovation noise  $\mathbf{e} \sim \mathcal{N}(0,1)$ , Signal-to-Noise Ratio 4.7 dB

## Numerical Example

- ▶ the first local model  $\tilde{A}_1 = A_1 K_1 C$  is not observable, which is a typical assumption in few PBSID algorithms.
- We run the proposed Algorithm with  $\mathcal{I}=50$  iterations and with the following *n*-selection pair  $(\alpha,\beta)$ , with n=3,

$$\alpha = \{(\epsilon, 1), (1, 1), (21, 1)\}, \ \beta = \{(\epsilon, 1), (2, 1), (21, 1)\}$$

▶ Best Fit Rate (BFR) and Variance Accounted For (VAF)

BFR	93.98 %
VAF	99.74 %

► The PBSID [van Wingerden et.al. 2009] approach failed for this example.

#### Conclusions

- Computationally efficient algorithm is proposed based on the idea of selections.
- The proposed algorithm avoids the curse of dimensionality.
- The algorithm is provenly consistent for a class of scheduling signals.
- An extension of the proposed approach with exogenous inputs signals is recently presented in LPVS 2019, Eindhoven, The Netherlands.

Thank you